



## Infectious Disease Practice

## Vaccination against measles-mumps-rubella and rates of non-targeted infectious disease hospitalisations: Nationwide register-based cohort studies in Denmark, Finland, Norway, and Sweden



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## SUMMARY

**Objectives:** To investigate if receipt of measles-mumps-rubella (MMR) vaccine following the third dose of diphtheria-tetanus-acellular pertussis (DTaP3) is associated with reduced rates of non-targeted infectious disease hospitalisations.

**Methods:** Register based cohort study following 1,397,027 children born in Denmark, Finland, Norway, and Sweden until 2 years of age. Rates of infectious disease hospitalisations with minimum one overnight stay according to time-varying vaccination status were compared using Cox proportional hazards regression analysis with age as the underlying timescale and including multiple covariates. Summary estimates were calculated using random-effects meta-analysis.

**Results:** Compared with DTaP3 and no MMR vaccine, MMR after DTaP3 was associated with reduced rates of infectious disease hospitalisations: aHR was 0.86 (0.83–0.89) in Denmark, 0.70 (0.64–0.75) in Finland, 0.71 (0.68–0.74) in Norway, and 0.71 (0.65–0.77) in Sweden: summary estimate was 0.75 (0.65 to 0.84). A beneficial association was also seen in a negative control exposure analysis (3 vs. 2 DTaP doses): summary estimate aHR was 0.81 (0.75–0.87).

**Conclusions:** Having MMR as the most recent vaccine was consistently associated with reduced rates of infectious disease hospitalisation. However, bias may account for at least some of the observed association. Randomised controlled trials are warranted to inform the optimal timing of MMR for both its specific and potential non-specific effects.

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## Introduction

Vaccines have been found to have non-specific effects (NSEs), affecting susceptibility towards other infections than the vaccine-targeted diseases.<sup>1</sup> The observed NSEs have varied depending on sex, the type of infection being studied, severity of disease, the type of vaccine being administered, and sequence of vaccinations. Live vaccines have often been associated with beneficial NSEs, which are most pronounced as long as the vaccines are the most recent vaccine administered.<sup>1</sup>

The initial observations were done in settings with high child mortality, where live vaccines were found to reduce child mortality more than what could be explained by the specific disease protection.<sup>2,3</sup> Studies from high-income countries, with low child mortality, have found similar patterns when looking at hospitalisations for non-targeted infectious diseases. Compared with having the non-live vaccine against diphtheria, tetanus, acellular pertussis, polio, and *Haemophilus influenzae* type b as the most recent vaccine, vaccination with the live combination vaccine against measles, mumps and rubella (MMR) has been associated with reduced rates of hospitalisations from non-targeted infections.<sup>4–8</sup> However, it is difficult to draw clear conclusions from these studies due to high risk of residual confounding.<sup>9</sup> Moreover, differences in settings and study protocols hamper comparisons of results.<sup>10–12</sup> Triangulation of results from multiple settings,<sup>13,14</sup> and employing identical analysis plans<sup>15</sup> have been proposed as methods to strengthen the causal deductions that can be made from observational studies.

The aim of this study was to investigate if receipt of MMR vaccine after the third dose of diphtheria-tetanus-acellular pertussis-containing vaccine (DTaP) was associated with lower rates of vaccine non-targeted infectious disease hospitalisation than receipt of three doses DTaP vaccine only, among children below 2 years of age born in Denmark, Finland, Norway, or Sweden, using national register data, similar analysis plans, and extensive control for potential confounders.

## Methods

The northern European countries Denmark, Finland, Norway, and Sweden (henceforward referred to as the Nordic countries) all have universal tax-funded health care, comparable socio-demographic characteristics, and extensive nationwide registries holding information on a multitude of health and sociodemographic information.<sup>16</sup> In all countries, the personal ID given upon birth or taking residency in the country makes linkage of data from the different registries possible.<sup>16</sup>

This register-based cohort study utilises the data collected within the Nordic collaboration “NONSENSE”. Due to current legislation, data was stored in each country separately. Description of settings, data sources, and harmonisation of data within NONSENSE is published elsewhere.<sup>17</sup>

This study included children born in the respective country from 1 January 2008 in Denmark and Norway, 1 July 2010 in Finland, and 1 January 2013 in Sweden until and including 31 December 2015 in all countries. The study period was based on availability of register data collected for the NONSENSE project, and pneumococcal conjugate vaccines (PCV) being used in the childhood immunisation programmes in all countries. DTaP and PCV were recommended at 3, 5, and 12 months of age in all countries. MMR vaccine was recommended at 15 months of age in Denmark and Norway, and 18 months of age in Sweden. In Finland, MMR was recommended together with DTaP at 12 months of age, but some children still received MMR after the third dose of DTaP. All vaccines within the Nordic childhood immunisation programmes are voluntary and administered free of charge. Individual-level information on

administered vaccines, including type and date of vaccination, is recorded in national vaccination registries.<sup>18–20</sup>

### Hospitalisations for infections

Hospital care for children is free of charge in Denmark and Norway.<sup>21,22</sup> A small patient fee up to an annual maximum amount may be charged for inpatient contacts in Finland and some regions in Sweden.<sup>23,24</sup> Individual-level information on all hospital contacts, including diagnoses, and dates of admission and discharge, is registered in nationwide patient registries.<sup>25–28</sup> Since 1997, diagnoses have been coded according to the International Classification of Diseases version 10 (ICD-10) in all countries.<sup>29</sup> The primary outcome was defined as inpatient contacts with overnight stays for any type of infection, including primary and secondary diagnoses ([sMaterial 1](#)), as this outcome has been found to occur at similar rates across the Nordic countries.<sup>30</sup> Secondary outcomes included inpatient contacts with at least two overnight stays (representing the more severe infections) and inpatient contacts with at least one overnight stay by type of infection, categorised as upper respiratory tract infections (URTI), lower respiratory tract infections (LRTI), gastrointestinal infections (GI) or other infections (OI) ([sMaterial 1](#)).

### Covariate assessments

From the nationwide registries, we included information on year and season of birth, sex, birth weight, mode of delivery, maternal smoking during pregnancy, singleton, child order, maternal age, maternal origin, household income quintile, single parenthood, maternal highest attained education, number of inpatient hospital contacts before 12 months of age, presence of chronic diseases, and receipt of other live or non-live vaccines (categorisation presented in [sMaterial 2](#)).

### Study design

We included children who had received the second dose of DTaP but neither the third dose of DTaP, first dose of MMR nor any other measles containing vaccine before 11 months of age ([sFigure 1](#)), to minimise bias related to reasons for non-vaccination. In the main analysis, we included children regardless of the number of registered PCV vaccinations.

Vaccination status was time varying and changed on the date of vaccination for each DTaP vaccine or MMR vaccine after baseline. Vaccination status was categorised as: 1) three doses of DTaP (DTaP3), 2) MMR given after three doses of DTaP (MMR-after-DTaP3), 3) concurrent MMR and DTaP3 vaccination (MMR-with-DTaP3), 4) MMR given after DTaP2 (MMR-after-DTaP2), and 5) DTaP3 given after MMR (DTaP3-after-MMR).

Inverse probability of treatment weights (IPTW) predicts the inverse probability of being exposed to different vaccination statuses as a function of the included covariates. We estimated the IPTW given the included covariates using multinomial logistic regression and truncated weights above the 99th percentile.<sup>31–33</sup> The IPTW were estimated in 14-day age intervals reflecting an age-dependent probability of vaccination given the covariates.

We limited follow-up to age intervals where a sufficient number of children had received MMR, i.e., excluding younger ages where only a few children had received MMR earlier than recommended. Baseline was defined based on visual inspection of vaccine exposure distribution ([sFigure 2](#)) and IPTW plots according to age (further described in [sMaterial 3](#), [sFigure 2–4](#)). Based on this evaluation, baseline was defined as 2 weeks prior to the age of recommended MMR vaccination in Denmark, Norway, and Sweden. In Finland, where MMR is recommended together with DTaP3, children who received MMR separately were generally older than the age of

recommended vaccination, and the baseline in Finland was thus set at 2 weeks after the age of recommended vaccination.

### Statistical analysis

The analyses were conducted in each country separately, using identical statistical coding in Stata 16 and/or 17. Children were followed from their date of vaccination (first of MMR or DTaP3), or from baseline whichever occurred last, and until 2 years of age, death, emigration, receipt of a fourth dose of DTaP, a second dose of MMR, or 31st December 2017, whichever occurred first (sFigure 1).

Infectious disease hospitalisations were included as recurrent events. Events that occurred within 14 days of a previous event were regarded as belonging to the same infectious disease episode. Therefore, the 13-day period after each event was censored and follow-up was restarted on day 14 after the previous event.

We first calculated crude rates of infectious disease hospitalisations as the number of events per 100 person-years. We used Cox proportional hazards regression model with age as the underlying timescale and repeated events (Andersen-Gill model<sup>34</sup>) to estimate the hazard ratios (HR) of infectious disease hospitalisations according to vaccination status and 95% confidence intervals (CI). We estimated; 1) unadjusted HRs; 2) covariate-adjusted HRs (aHR) and 3) IPTW HRs using time-varying weights. The IPTW model was further adjusted for the included covariates to account for remaining covariate imbalance after weighting.

First, the analyses were performed for hospitalisations with minimum one and two overnight stays, respectively, including all types of infections. Second, the analysis for infectious disease hospitalisations with minimum one overnight stay was performed by type of infection. All analyses were performed for all children combined and by sex. We used the Wald test for interaction to identify potential sex differential effects.

For the primary outcome in the covariate-adjusted model, the proportional hazards assumption was tested using Schoenfeld residuals,<sup>35</sup> if violations were observed between exposure groups, we estimated the HR in 8-week follow-up intervals.

Summary estimates across all countries were calculated using the DerSimonian-Laird method for random-effects meta-analysis accounting for between study heterogeneity.<sup>36</sup>

### Sensitivity and subgroup analyses

Children who leave their country of residence temporarily (up to 12 months) are not required to be registered as emigrants, causing loss to follow-up without the possibility to censor them. Children with a parent born abroad may be more likely to leave the country for longer periods of time. MMR is recommended to be given prior to travelling abroad (specific recommendations vary between countries), which could lead to an underestimation of events among children who have received MMR. This could bias the results towards a beneficial effect of MMR. Thus, we performed a subgroup analysis restricted to children with two native-born parents.

Families are generally advised to have the MMR-vaccination postponed if the child has fever. Thus, children will tend to be free from illness at the time of vaccination, introducing healthy vaccinee bias.<sup>37</sup> At the same time, MMR can give transient fever. We therefore conducted a sensitivity analysis excluding the 14 days after vaccination with MMR from follow-up.

In Finland, children are recommended annual seasonal influenza vaccination from 6 months of age; we investigated if receipt of influenza vaccine affected the results by censoring children upon influenza vaccination.

Missing PCV and rota virus vaccine (RV) vaccinations could indicate vaccine hesitancy, which could also apply to MMR. We conducted a subgroup analysis excluding children who had not received

two doses of PCV, and RV as recommended in each country, before 11 months of age.

We explored the presence of unmeasured bias attributable to not receiving vaccines as recommended by investigating the rate of infectious disease hospitalisations with minimum one overnight stay among children who had received the third dose of DTaP and no MMR, compared with children who had received two doses of DTaP, as negative control exposure. Post hoc, we further performed this negative control analysis for the different types of infections. In this analysis, we followed children from 11 months of age until 15 months of age, death, migration, or receipt of MMR, whichever came first.

Finally, we calculated the G-value for the strength required by an unmeasured confounder to return the observed aHR of infectious disease hospitalisations in the main analysis to the null.<sup>38</sup>

### Ethical approvals

The study was approved by the Regional Ethics Committee, South-East, in Norway and by the Regional Ethical Review Board, Stockholm, Sweden. Ethical approval is not required for registry-based studies in Denmark or Finland, but the study was approved by the Danish Data Protection Agency and the Institutional Review Board of the Finnish Institute for Health and Welfare.

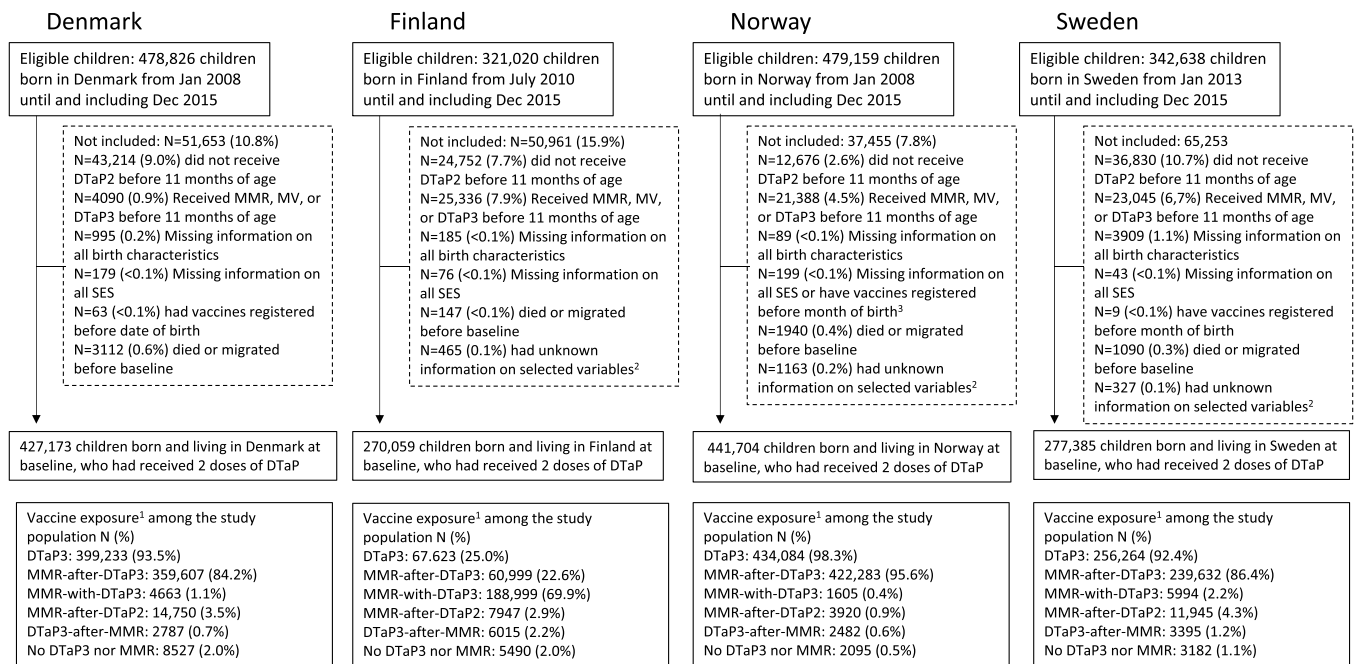
### Results

A total of 1,621,643 children were born in the countries during the respective study periods. After exclusions, primarily of children who had received measles-containing vaccines or not received 2 doses of DTaP before 11 months of age, 1,397,027 children were included (Fig. 1).

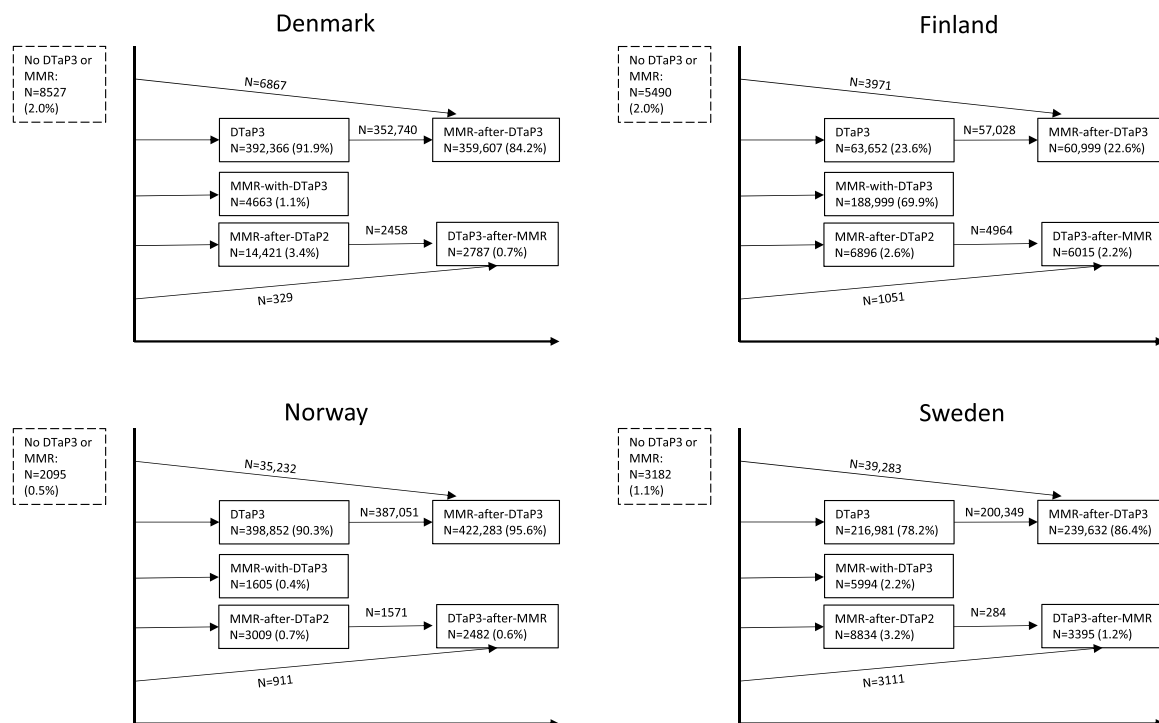
In Denmark, Norway and Sweden, most children (>92%) followed the recommended vaccination sequence and received 3 doses of DTaP before any MMR (Fig. 1). In Finland, 25% received DTaP3 before any MMR, whereas most children got the MMR and DTaP3 together as recommended (Fig. 1). The age at vaccination with MMR after DTaP3 varied the most in Finland with an interquartile range (IQR) of 68 days, followed by 54 days in Denmark, compared with an IQR of 32 days in Norway and 27 days in Sweden (sTable 1). Among children with DTaP3, the proportion who subsequently received MMR before 2 years of age was lower in Denmark (90.1%) and Finland (90.7%) compared with Sweden (93.5%) and Norway (97.3%) (Fig. 2, sFigure 2 and calculated from number of children in each exposure group presented in Fig. 1).

The children in the vaccination groups MMR-with-DTaP3, MMR-after-DTaP2, and DTaP3-after-MMR represent small subgroups of the study population (up to 3.4%) who do not follow the recommended vaccination sequence, except for the MMR-with-DTaP3 group in Finland (Fig. 2). The median age at entering the respective vaccination groups (sTable 1) indicated differences across countries concerning whether the children in these subgroups received the vaccines prior to, according to, or later than recommended. The present analysis focused on the effect of having MMR after DTaP3 compared with DTaP3 and no MMR.

Compared to children who had received MMR after DTaP3 one month after recommended MMR vaccination, children who had not yet received MMR were less likely to be firstborn and more likely to come from families with low household income and low maternal education in Denmark, Norway, and Sweden. This pattern was not seen in Finland (Table 1). Overall, the associations between covariates and vaccination status at two years of age were similar to the associations at one month after recommended vaccination (sTable2).



**Fig. 1.** Flowchart for study population participation including number of children who enter each vaccination group. Abbreviations: MMR: Measles, Mumps, Rubella vaccine; MV: measles containing vaccines; DTaP3: received 3 doses of diphtheria, tetanus, acellular pertussis, polio, and Haemophilus Influenzae type b vaccine; SES: Socioeconomic status; MMR-after-DTaP3: received MMR after DTaP3; MMR-with-DTaP: concurrent MMR and DTaP3 vaccination; MMR-after-DTaP2: MMR after second dose of DTaP; DTaP3-after-MMR: third dose of DTaP received after MMR vaccination. <sup>1</sup>Vaccination statuses of the children in the study population after receipt of the second dose of DTaP (inclusion criteria). The proportion of children in the DTaP3 and MMR-after-DTaP2 vaccination group does not reflect the number of children that contribute with follow-up in these groups as some may have received a subsequent MMR or DTaP3 vaccine before start of follow-up. The groups are not exclusive as the child can contribute to multiple vaccination groups before 2 years of age. <sup>2</sup>we excluded children that had unknown information on a variable where less than 2 per thousand had missing information due to low numbers in these strata. <sup>3</sup>In Norway only months of birth was available. We assigned an exact date of birth to each child as a random integer within the month of birth.



**Fig. 2.** Changes in vaccination status during follow-up. Abbreviations: DTaP3: received 3 doses of diphtheria, tetanus, acellular pertussis, polio, and Haemophilus Influenzae type b vaccine; MMR-after-DTaP3: received Measles, Mumps, rubella vaccine (MMR) after DTaP3; MMR-with-DTaP3: concurrent MMR and DTaP3 vaccination; MMR-after-DTaP2: MMR after second dose of DTaP; DTaP3-after-MMR: third dose of DTaP received after MMR vaccination. Number of children that belong to the different vaccination groups during follow-up (numbers within the boxes), proportions calculated with number of children included in each of the countries as the denominator: 427,173 in Denmark, 270,059 in Finland, 441,704 in Norway, and 277,385 in Sweden (Fig. 1). Arrows going from the y-axis indicate persons belonging to that vaccination group at start of follow-up (date of DTaP3 or MMR or baseline age, see methods). Arrows between vaccination groups indicate number of children moving from one vaccination group to another vaccination group during follow-up. Note that numbers and proportions do not sum to the total number, because a child can belong to several vaccination groups during follow-up.

**Table 1**  
Characteristics of children exposed to DTaP3 or MMR after DTaP3 at 1 month after age of recommended MMR vaccination<sup>a</sup> in Denmark, Finland, Norway, and Sweden.

	Denmark			Finland			Norway			Sweden		
	DTaP3	MMR after DTaP3		DTaP3	MMR after DTaP3		DTaP3	MMR after DTaP3		DTaP3	MMR after DTaP3	
	169,638	215,597		46,557	12,543		100,600	329,202		63,494	192,035	
Study population												
Sex												
Male	87,823	108,907	51.8%	23,588	6389	50.9%	52,482	168,184	51.1%	33,135	98,463	51.3%
Female	81,815	106,690	48.2%	22,969	6154	49.1%	48,118	161,018	48.9%	30,359	93,572	48.7%
Year of birth												
2008	23,672	25,982	14.0%	0	0	0.0%	17,064	36,876	11.2%	0	0	0.0%
2009	22,601	26,695	13.3%	0	0	0.0%	15,690	39,840	12.1%	0	0	0.0%
2010	21,801	29,418	12.9%	6876	1980	15.8%	14,131	39,840	12.1%	0	0	0.0%
2011	22,073	26,011	13.0%	10,133	3231	21.8%	13,807	40,304	13.7%	0	0	0.0%
2012	21,789	25,594	12.8%	10,036	2552	20.3%	12,291	41,288	12.2%	0	0	0.0%
2013	20,117	24,850	11.9%	7749	1846	16.6%	10,968	41,690	12.7%	22,526	61,352	31.9%
2014	17,893	27,243	10.5%	6276	1510	12.0%	8486	43,438	13.2%	21,302	63,600	33.1%
2015	19,692	29,804	11.6%	5487	1424	11.4%	8163	44,584	13.5%	19,666	67,083	34.9%
Season of Birth												
Winter (December to February)	38,888	51,611	22.9%	11,323	3155	25.2%	19,393	80,549	24.5%	16,468	43,116	22.5%
Spring (March to May)	41,247	53,327	24.3%	9909	2602	20.7%	28,477	82,589	25.1%	15,436	51,732	26.9%
Summer (June to August)	45,366	58,643	26.7%	12,190	3222	25.7%	27,304	87,980	26.7%	18,483	48,645	25.3%
Autumn (September to November)	44,137	52,016	26.0%	13,135	3564	28.4%	25,426	78,084	23.7%	13,107	48,542	25.3%
Birth weight (grams)												
< 2000	3250	3455	1.9%	776	202	1.6%	2003	5301	1.6%	813	1864	1.0%
2001–2500	5598	7059	3.3%	1240	303	2.4%	3059	9379	2.8%	1737	4917	2.6%
2501–3000	20,749	27,550	12.2%	5204	1399	11.2%	11,498	37,689	11.4%	7160	21,743	11.3%
3001–3500	53,689	71,066	31.6%	15,083	4134	33.0%	31,781	106,235	32.3%	20,494	61,957	32.3%
3501–4000	56,384	71,518	33.2%	16,522	4481	35.7%	34,651	114,077	34.7%	21,500	66,471	34.6%
4001–4500	22,896	27,316	13.5%	6621	1736	13.8%	14,482	46,662	14.2%	9520	28,527	14.9%
> 4500	4673	5284	2.8%	1111	288	2.3%	3126	9859	3.0%	2270	6556	3.4%
Unknown	2399	2349	1.4%	0	0	0.0%	0	0	0.0%	0	0	0.0%
Preterm (born before 37 full weeks of gestation)												
Not preterm	157,034	201,064	92.6%	43,913	11,872	94.7%	93,415	308,167	93.6%	60,040	182,683	95.1%
Preterm	10,790	12,885	6.4%	2644	671	5.3%	6632	19,486	5.9%	3454	9352	4.9%
Unknown	1814	1648	1.1%	0	0	0.0%	553	1549	0.5%	0	0	0.0%
Delivered by caesarean section												
Not delivered by caesarean section	132,077	167,416	77.9%	38,600	10,417	83.1%	83,055	274,523	83.4%	52,204	158,379	82.2%
Delivered by caesarean section	37,561	48,181	22.1%	7957	2126	16.9%	17,545	54,679	16.6%	11,290	33,656	17.5%
Unknown	0	0	0.0%	0	0	0.0%	0	0	0.0%	0	0	0.0%
Maternal smoking or snuff during pregnancy												
No smoking or snuff during pregnancy	144,745	189,727	85.3%	38,506	10,539	84.0%	73,232	251,367	76.4%	56,778	175,571	91.4%
Smoking or snuff during pregnancy	20,802	21,627	12.3%	7009	1758	14.0%	10,387	26,836	8.2%	4337	9266	4.8%
Unknown	4091	4243	2.4%	1042	246	2.0%	16,981	50,999	15.5%	2379	7198	3.7%
Singleton												
No	7450	8677	4.4%	1350	304	2.4%	3751	10,102	3.1%	1986	5149	2.7%
Yes	162,188	206,920	95.6%	45,207	12,239	97.6%	96,849	319,100	96.9%	61,508	186,886	97.3%
Unknown	0	0	0.0%	0	0	0.0%	0	0	0.0%	0	0	0.0%
Child order												
Firstborn	62,559	112,345	36.9%	18,292	5677	45.3%	36,451	146,584	44.5%	23,073	85,077	44.3%
1 older sibling	69,486	72,136	41.0%	16,046	4170	33.2%	39,248	116,945	35.5%	25,120	71,524	37.2%
2 older siblings	27,247	23,033	16.1%	7304	1759	14.0%	17,334	48,017	14.6%	10,088	25,155	13.1%
3 or more older siblings	9032	6116	5.3%	4915	937	7.5%	7567	17,656	5.4%	5213	10,279	5.4%
Unknown	1314	1967	0.8%	0	0	0.0%	0	0	0.0%	0	0	0.0%
Maternal age at delivery												
< 20 years	4161	4464	2.5%	1828	440	3.5%	3884	10,133	3.1%	1719	3807	2.0%
21–25 years	22,723	30,613	13.4%	8103	2074	16.5%	17,694	56,545	17.2%	11,385	31,661	16.5%
26–30 years	55,477	78,531	32.7%	15,333	4245	33.8%	33,154	112,720	34.2%	20,664	64,203	33.4%

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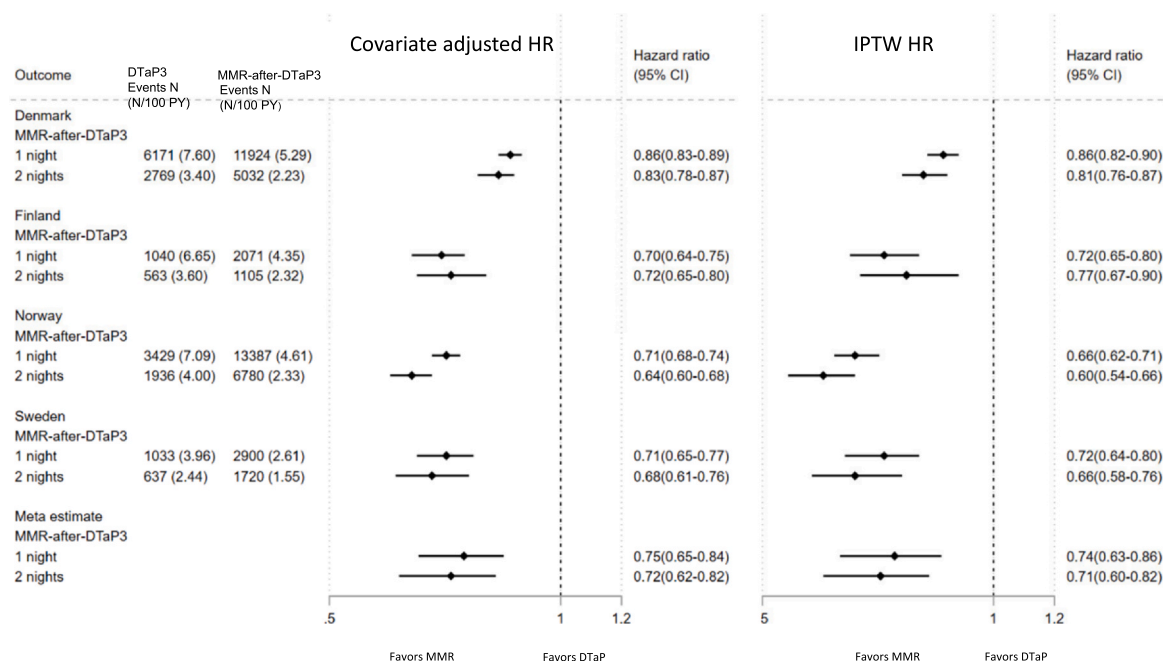


Table 1 (continued)

	Denmark			Finland			Norway			Sweden		
	DTaP3	MMR after DTaP3	DTaP3	MMR after DTaP3	DTaP3	MMR after DTaP3	DTaP3	MMR after DTaP3	DTaP3	MMR after DTaP3	DTaP3	MMR after DTaP3
Study population	169,638	215,597	46,557	12,543	100,600	329,202	63,494	192,035	192,035	60,573	31,5%	16.6%
31–35 years	58,462	70,058	14,071	3942	30,447	100,643	18,976	60,573	60,573	31,791	16.6%	0.0%
36 years and older	28,815	31,931	7222	1842	15,421	49,161	10,750	31,791	31,791	0	0.0%	0.0%
Unknown	0	0	0	0	0	0	0	0	0	0	0.0%	0.0%
Maternal origin												
Born in-country	140,522	183,115	42,998	11,506	76,839	256,075	45,360	146,411	146,411	45,624	76.2%	23.8%
Born abroad	29,116	32,482	3559	1037	23,761	73,127	18,134	45,624	45,624	0	0.0%	0.0%
Unknown	0	0	0	0	0	0	0	0	0	0	0.0%	0.0%
Household income quintile at birth												
First (lowest)	32,153	32,655	8968	2261	21,671	58,235	13,354	31,476	31,476	13,354	21.0%	16.4%
Second	32,865	39,235	9773	2444	21,237	64,808	13,994	38,416	38,416	13,994	22.0%	20.0%
Third	32,663	42,174	9825	2559	20,422	67,484	13,126	41,051	41,051	13,126	20.7%	21.4%
Fourth	31,545	44,940	9415	2697	19,030	69,277	12,116	41,308	41,308	12,116	19.1%	21.5%
Fifth (highest)	29,893	46,359	8399	2533	18,240	69,398	10,904	39,784	39,784	10,904	17.2%	20.7%
Unknown	10,519	10,234	177	49	0	0	0	0	0	0	0.0%	0.0%
Single parenthood												
Lives with 2 parents	155,052	200,696	42,974	11,650	89,480	300,289	55,367	173,889	173,889	55,367	87.2%	90.6%
Lives with single parent	14,073	14,430	3350	840	11,120	28,913	8127	18,146	18,146	8127	12.8%	9.4%
Unknown	513	471	233	53	0	0	0	0	0	0	0.0%	0.0%
Maternal highest attained education <sup>b</sup>												
Low education	26,816	25,998	5682	1368	21,088	55,742	9661	20,013	20,013	9661	15.2%	10.4%
Medium education	57,764	73,008	19,494	5008	27,498	88,081	18,796	58,647	58,647	18,796	29.6%	30.5%
High education	81,206	112,414	21,381	6167	48,173	175,069	20,353	73,203	73,203	20,353	32.1%	38.1%
Unknown	3852	4177	0	0	3841	10,310	14,684	40,172	40,172	14,684	23.1%	20.9%
Chronic diseases <sup>c</sup>												
No chronic diseases	163,870	209,110	43,238	11,759	93,408	309,630	61,187	186,335	186,335	61,187	96.4%	97.0%
Chronic diseases	5768	6487	3319	784	7192	19,572	2307	5700	5700	2307	3.6%	3.0%
Received other live vaccines												
Yes	691	1325	42,310	12,103	25,652	99,080	0	0	0	0	0.0%	0.0%
Unknown <sup>d</sup>	168,947	214,272	42,47	440	74,948	230,122	63,494	192,035	192,035	63,494	100.0%	100.0%
Received other non-live vaccines <sup>e</sup>												
Yes	4183	9422	7567	3524	27,663	79,529	39	98	98	27,663	0.1%	0.1%
Unknown <sup>d</sup>	165,455	206,175	38,990	9019	72,937	249,673	63,455	191,937	191,937	63,455	99.9%	99.9%
Number of hospital contacts before 12 months of age												
0	98,581	126,764	35,142	9581	75,001	249,862	50,639	156,559	156,559	50,639	79.8%	81.5%
1	47,048	60,882	8568	2305	18,922	62,201	9784	28,141	28,141	9784	15.4%	14.7%
2	14,684	17,923	1787	421	4005	11,388	1989	5168	5168	1989	3.1%	2.7%
3	5054	5542	565	117	1332	3112	545	1284	1284	545	0.9%	0.7%
4 or more	4271	4486	495	119	1340	2639	537	883	883	537	0.8%	0.5%

Abbreviations: DTaP3 = third dose of the non-live vaccine against diphtheria, tetanus, acellular pertussis, polio, and Haemophilus influenzae type b; MMR= vaccination against measles, mumps, and rubella.

<sup>a</sup> One month after age of recommended vaccination with MMR reflects the date the child turns 13 months of age in Finland, 16 months of age in Denmark and Norway, and 19 months of age in Sweden.<sup>b</sup> Highest attained education classified according to ISCED2011.<sup>c</sup> Chronic diseases are defined according to the definition provided by Kristensen et al.<sup>46</sup><sup>d</sup> Reporting of vaccines not belonging to the immunisation programme was not mandatory in all countries, thus lacking registration of other vaccines may not reflect that no other vaccines was given.<sup>e</sup> PCV was recommended to all children throughout the study period. PCV was not included in other non-live vaccines.



**Fig. 3.** Hazard ratios of infectious disease hospitalisations with a minimum of 1 or 2 overnight stays among children with MMR after three doses of DTaP compared with 3 doses of DTaP without MMR, by country and combined in summary estimate. Abbreviations: DTaP3: Received 3 doses of diphtheria, tetanus, acellular pertussis, polio, and Haemophilus Influenzae type b vaccine; MMR-after-DTaP3: Received measles, mumps, rubella vaccine after DTaP3; HR: Hazard Ratio; IPTW: Inverse probability of treatment weighted Estimated using an extended Cox regression with age as the underlying time scale, vaccination status included as time-varying exposure and infectious disease hospitalisations included as recurrent events. Summary estimate is calculated using DerSimonian-Laird method for random-effects meta-analysis. Covariates included in both the adjusted and weighted model: Year and season of birth, sex, birth weight, mode of delivery, maternal smoking during pregnancy, singleton, child order, maternal age, maternal origin, household income quintile, single parenthood, maternal highest attained education, number of inpatient hospital contacts before 12 months of age, chronic diseases, and receipt of other live or non-live vaccines.

### Infectious disease hospitalisations

In all countries, children with MMR after DTaP3 had a lower rate of infectious disease hospitalisations with overnight stays compared with children not vaccinated with MMR (the DTaP3 exposure group): the unadjusted HR was 0.81 (0.78 to 0.84) in Denmark, 0.68 (0.63 to 0.73) in Finland, 0.64 (0.61 to 0.67) in Norway, and 0.66 (0.61 to 0.71) in Sweden (sTable 3). The aHRs were 0.86 (0.83–0.89) in Denmark, 0.70 (0.64–0.75) in Finland, 0.71 (0.68–0.74) in Norway, and 0.71 (0.65–0.77) in Sweden (sTable 3, Fig. 3). The summary estimate for the aHR across countries was 0.75 (0.65 to 0.84) (Fig. 3). The summary estimate was 0.73 (0.62 to 0.84) for boys and 0.78 (0.69 to 0.87) for girls (sFigure 5). The summary estimate for hospitalisations with at least two overnight stays was 0.72 (0.62 to 0.82) (Fig. 3). The IPTW HRs were similar to the aHRs (Fig. 3, sTable 3, sTable 4). There were non-proportional hazards in Norway with the highest aHR for infectious disease hospitalisations with minimum one overnight stay of 0.83 (0.77 to 0.90) observed in the first 8 weeks of follow-up (sTable 5).

### Type of infection

The hazard of infectious disease hospitalisations among children exposed to MMR after DTaP3 compared with DTaP3 was lower for all types of infections (Fig. 4, sTable 6). The association was strongest for LRTI in Denmark, strongest for GI in Finland and OI in Sweden (Fig. 4). In Norway the results differed between the two models. The association was strongest for LRTI in the covariate adjusted model and for URTI in the IPTW model (Fig. 4). The associations by type of infection were similar for boys and girls (sFigure 6).

### Sensitivity- and subgroup analyses

Restricting the analyses to children of parents born in the respective countries (sTable 7), censoring 14 days after receipt of MMR (sTable 8), censoring upon influenza vaccination in Finland

(sTable 9), or restricting the study population to children who had received two doses of PCV before 11 months of age, and RV as recommended, (sTable 10) did not substantially change the results.

Having received three doses of DTaP (DTaP3) compared with two (DTaP2) was associated with reduced rates of infectious disease hospitalisations in all countries: aHR was 0.83 (0.79 to 0.86) in Denmark, 0.89 (0.80 to 1.00) in Finland, 0.74 (0.70 to 0.79) in Norway, and 0.81 (0.73 to 0.89) in Sweden, yielding a summary estimate across countries of 0.81 (0.75 to 0.87) (Fig. 5, sTable 11). The summary estimates across countries were similar for the different types of infections.

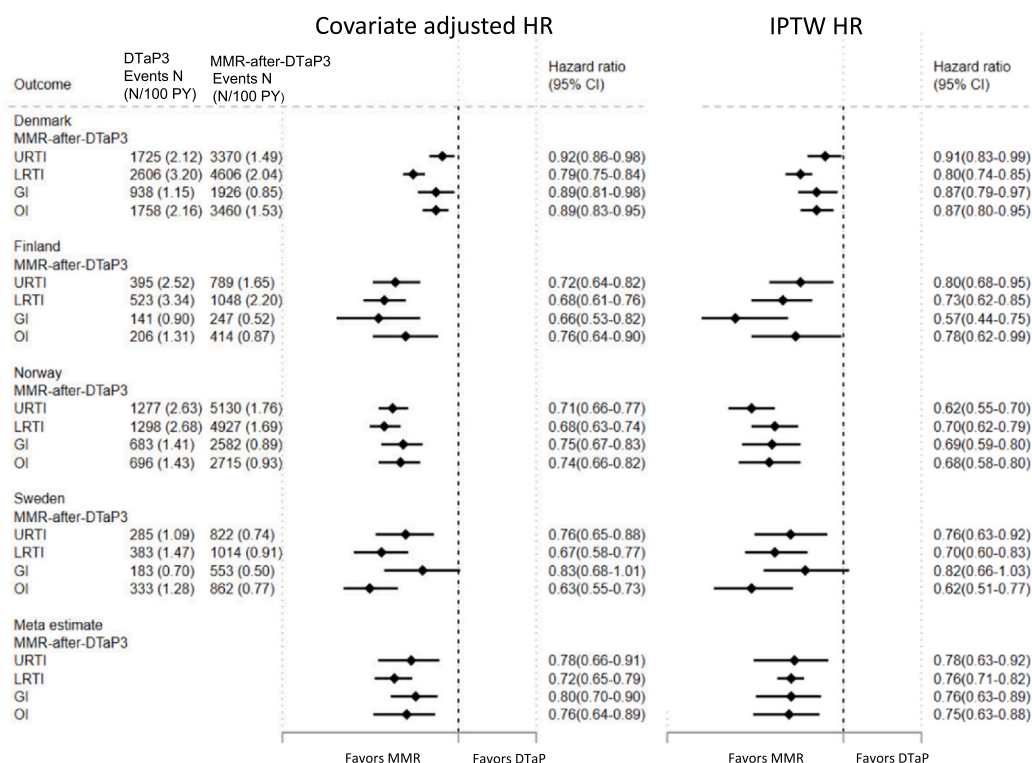
G-values for the required strength of confounding to return the observed aHR in the main analysis to the null was 1.6 in Denmark, 2.12 in Finland and Sweden, and 2.40 in Norway if the confounding factor is present in 100% of children with DTaP3 (sTable 12).

## Discussion

Receipt of MMR after DTaP3 was consistently associated with lower rates of infectious disease hospitalisations compared with not having received MMR after DTaP3. The reduction was smaller in Denmark (14%) than in the other Nordic countries (approximately 30%). Similar protective associations were seen after receipt of DTaP3 when compared to DTaP2. We did not find any consistent differences in the results by sex, duration of hospitalisation, or type of infection across countries and statistical models.

### Strengths and limitations

The study was based on population-based registers that contained information on exposures, outcomes, and potential confounders. Further strengths pertain to the comparable data structure across countries by use of a common data model,<sup>17</sup> and previous



**Fig. 4.** Covariate adjusted, and inverse probability of treatment weighted (IPTW) Hazard Ratio of infectious disease hospitalisations with minimum 1 overnight stay for different types of infections for children with MMR after three doses of DTaP compared with 3 doses of DTaP without MMR, by country and combined in meta estimate. Abbreviations: DTaP3: received 3 doses of diphtheria, tetanus, acellular pertussis, polio, and Haemophilus Influenzae type b vaccine; MMR-after-DTaP3: received Measles, Mumps, rubella vaccine after DTaP3; HR: Hazard Ratio; IPTW: Inverse probability of treatment weighted; URTI: upper respiratory tract infections; LRTI: Lower respiratory tract infections; GI: Gastrointestinal infections; OI: Other infections. Estimated using an extended Cox regression with age as the underlying time scale, vaccination status included as time varying exposure and infectious disease hospitalisations included as recurrent events. Summary estimate is calculated using DerSimonian-Laird method for random-effects meta-analysis. Covariates included in both the adjusted and weighted model: Year and season of birth, sex, birth weight, mode of delivery, maternal smoking during pregnancy, singleton, child order, maternal age, maternal origin, household income quintile, single parenthood, maternal highest attained education, number of inpatient hospital contacts before 12 months of age, chronic diseases, and receipt of other live or non-live vaccines.

investigations of the outcomes to ensure comparability of measures across countries.<sup>30</sup>

However, observational studies such as ours are limited by an inherent risk of residual confounding. To reduce the risk of bias attributable to complete non-vaccination, we restricted the study population to children who had followed the vaccination programme for DTaP and MMR until 11 months of age. We used age as the underlying timescale, to ensure complete adjustment for age. We furthermore adjusted for a range of potential confounders, which attenuated the effect estimates in all countries to some extent. The IPTW model is based on the time-varying function of the covariates, which may offer better adjustment for reasons for delayed vaccination, but the results were quite similar in both the covariate-adjusted and the IPTW models. Noteworthy, estimated G-values indicate that unmeasured confounders must be associated with a 60–140% greater risk of infectious disease hospitalisations if present among 100% of children with DTaP3 to return the observed aHR to 1 – thus unmeasured confounders must represent stronger predictors of both vaccination status and risk of infectious disease hospitalisations than the combined covariates included in our analyses to return the observed aHR to 1.

#### Country-specific considerations

In all countries, the observed association may be biased towards a beneficial effect of receiving MMR if reasons for not receiving MMR as recommended are connected to an increased risk of infectious diseases. The lower MMR uptake with more delays in Denmark has

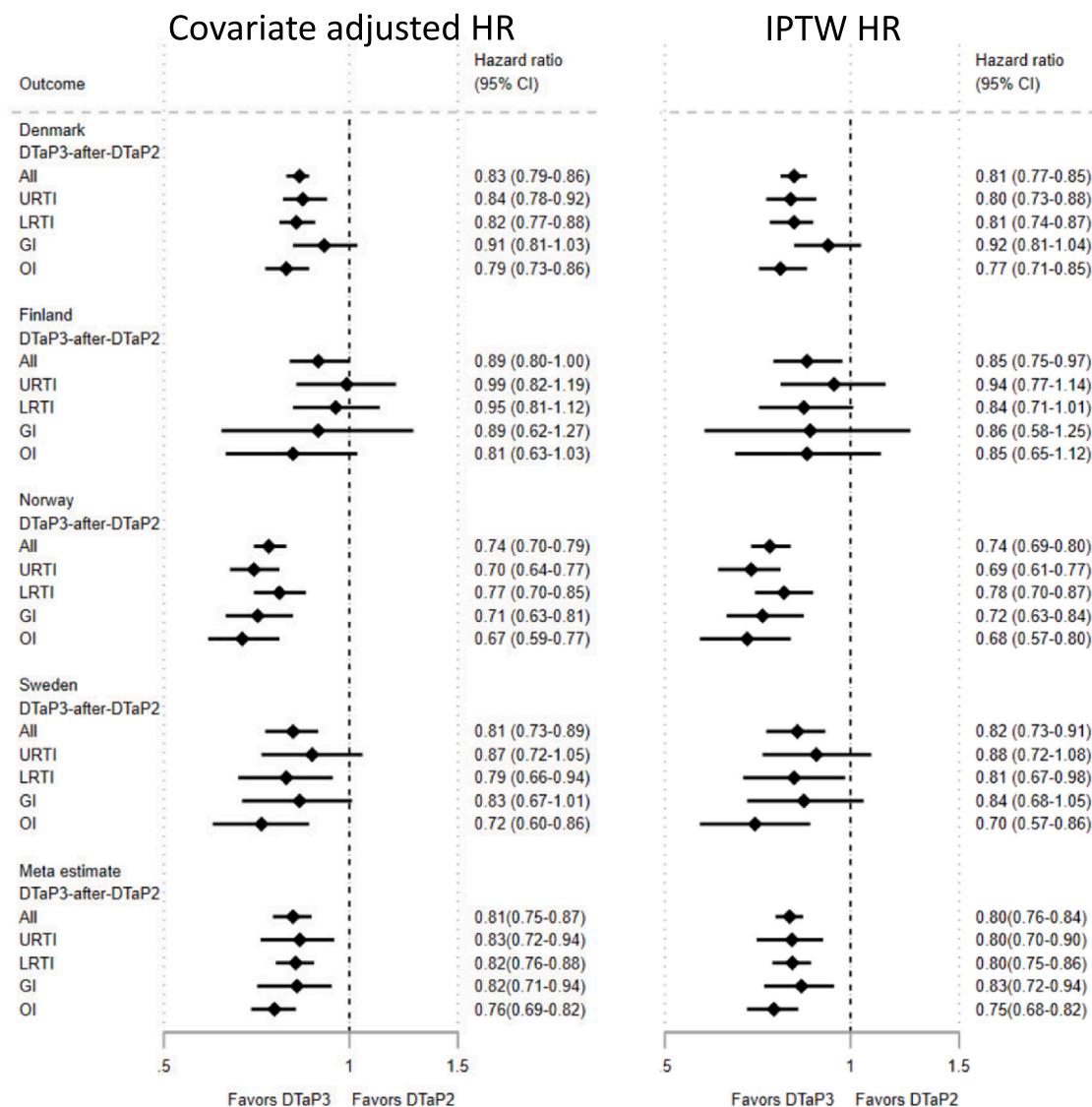
previously been ascribed to parents forgetting appointments or re-scheduling due to busy lives.<sup>39</sup> These more random delays may increase comparability of children who have received MMR vaccine and those who have not in Denmark. In contrast, in the other Nordic countries, particularly in Norway, with high and steep MMR uptake, the few children who remain MMR unvaccinated probably represent a more selected group, which may result in a greater bias towards a beneficial effect of MMR. Bias must be suspected to account for at least some of the observed association in all countries. In general, it has been observed that the higher and steeper the MMR uptake, the more beneficial the estimated effect (see also below). Thus, the greater beneficial association observed in the countries other than Denmark, likely reflects a greater degree of bias.

#### Negative control exposure

In the main analysis the reduction in hospitalisation following MMR vaccination was 25% (16–35%) across countries. However, the negative control analysis of DTaP3 vs DTaP2 also yielded a reduction in the rates of infectious disease hospitalisations (19% (95% CI: 13–25%)) across countries, indicating that bias related to not following the recommended vaccination schedule accounts for at least some of the observed association.

Tielemans et al.<sup>8</sup> also used a negative control exposure. In their main analysis, they found MMR+Meningococcal C vaccination compared with DTaP4 to be associated with a 38% (33–43%) lower rate of admission to hospital for infection. In their negative control





**Fig. 5.** Negative control exposure analysis: Hazard ratio of infectious disease hospitalisation when exposed to three doses of DTaP compared with two doses of DTaP for and hospitalisations combined and by type of infection, by country and combined in meta estimate. Abbreviations: DTaP2: received 2 doses of diphtheria, tetanus, acellular pertussis, polio, and Haemophilus Influenzae type b vaccine; DTaP3: Received 3 doses of DTaP; HR: Hazard Ratio; IPTW: Inverse probability of treatment weighted; All: all types of infections; URTI: upper respiratory tract infections; LRTI: Lower respiratory tract infections; GI: Gastrointestinal infections; OI: Other infections. Estimated using an extended Cox regression with age as the underlying time scale, vaccination status included as time-varying exposure and infectious disease hospitalisations included as recurrent events. Infectious disease hospitalisations that occur within 14 days from a previous infectious disease hospitalisation are regarded to belong to the same infectious disease episode. Children were followed from 11 months of age until 15 months of age, death, migration, or receipt of MMR, whichever occurred first. Adjusted for year and season of birth, sex, birth weight, mode of delivery, maternal smoking during pregnancy, singleton, child order, maternal age, maternal origin, household income quintile, single parenthood, maternal highest attained education, number of inpatient hospital contacts before 11 months of age, chronic diseases and receipt of other live or non-live vaccines.

analysis, having 4 doses of DTaP vs. DTaP3 was associated with a 31% (24–37%) lower rate of infectious disease hospitalisations.

We observed a more beneficial association in our main analysis than in the negative control analysis in all countries except Denmark, which is noteworthy, given that the uptake of DTaP3 was steeper, and the coverage was higher than for MMR (Material 4), indicating a bigger potential for healthy vaccinee bias in the negative control analysis. However, there may be other differences in the underlying bias structures in the two analyses and therefore any comparison of results should be interpreted with caution.

#### Comparison with other studies

Previous cohort studies have found lower rates of infectious disease hospitalisation when MMR as most recent vaccine is compared with DTaP,<sup>4–7</sup> also when MMR was given together with the

non-live vaccine against meningitis C.<sup>8</sup> The greatest protective associations were observed in countries with high and steep MMR uptake, and thus also an assumed larger degree of healthy vaccinee bias,<sup>5,7,8</sup> similar to what was observed in our study. Two previous studies have been conducted using Danish registry data, with divergent conclusions.<sup>5,7</sup> The first study found an aHR of 0.86 (0.84 to 0.88),<sup>5</sup> similar to the Danish estimates in our study (aHR: 0.86 (0.83 to 0.89)), ascribing it partly to an effect of MMR on hospital admissions for infections. The other study reported an aHR of 0.93 (0.92 to 0.94), but ascribed the association to residual confounding.<sup>7</sup> The latter study, however, did not make the same restrictions to the study population and followed children longer, up to 5 years of age, which may have contributed to the divergent conclusions.<sup>11</sup>

The beneficial effect of having MMR after DTaP was not found in a self-controlled case series analysis,<sup>12</sup> and was less clear<sup>40</sup> or weaker<sup>41</sup> when MMR was co-administered with DTaP.

Previous studies have found beneficial NSE of measles-containing vaccines to be most pronounced for LRTI.<sup>5,6,42,43</sup> This was not so clear in this study, and estimates were inconsistent between statistical models and countries. We have previously found variation across countries in the rates of infectious disease hospitalisations for the different types of infections.<sup>30</sup> Differences in coding practices may affect the results on type of infection differently in different countries. Furthermore, some of the previous studies that showed beneficial NSEs of MMR for respiratory tract infections were conducted among children not offered PCV. In all Nordic countries, PCV is offered together with DTaP, which may make the potential NSE of MMR vaccine less pronounced.

NSEs have also been found to vary by sex in studies from low-income settings.<sup>1</sup> We did not find consistent patterns of sex differential effects, in line with other studies from high income countries.<sup>5,6,8</sup>

### Immunological mechanisms

Trained innate immunity has been proposed as a concept and mechanism for NSEs of vaccines.<sup>44–46</sup> This involves long-term functional reprogramming of innate immune cells by stimulation with eg a live vaccine, leading to an altered innate immune response at future stimulation.<sup>44,45</sup> Most immunological studies have focused on the live Bacille Calmette-Guérin (BCG) vaccine against tuberculosis, but recently, transcriptional and functional alterations consistent with introduction of trained immunity in  $\gamma\delta$  T cells following MMR vaccination have been observed.<sup>47</sup>

### Conclusion and perspectives

In all four countries, we observed a lower rate of infectious disease hospitalisations among children who had MMR as their most recent vaccine compared with children who had not yet received MMR. Despite careful considerations of study design and control for numerous potential confounders, we cannot exclude residual confounding. This was illustrated by the protective association also seen for the control exposure (DTaP3). To further explore NSEs of MMR in high-income settings, randomised trials are warranted. Such randomised trials could be designed to examine the optimal age of MMR vaccination, with respect to both measles control and the occurrence of non-targeted infections.

### Transparency

The guarantor affirms that the manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as originally planned have been explained.

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### Contributors

LG, SM, HE, IL, HN, BF, AAP, LT, CSB, and SS conceptualized the study. All authors directed the analyses, which were carried out by LG with supervision of SM. LG, IL, HE, and ML managed data curation, undertook the country specific coding, and produced the country-specific results. LG, HE, IL, HN, BF, AAP, LT, and SS obtained the data to be included in this study. All authors contributed to the discussion and interpretation of the results. LG drafted the first

version of the manuscript. SM, HE, IL, HN, BF, ML, AAP, LT, CSB, and SS critically revised the draft. LG produced the visualizations. LG, IL, HN, BF, ML, AAP, LT, CSB, and SS obtained the funding for the present project. All authors approved the final version for submission. LG is the guarantor. The corresponding author attests that all listed authors meet authorship criteria and that no others meeting the criteria have been omitted. All authors had full access to all of the statistical reports and tables in the study and can take responsibility for the integrity of the data and the accuracy of the data analysis.

### Data availability

Due to data protection rules, we are not allowed to share the individual-level data, but other researchers fulfilling the requirements could obtain similar data from the register controllers.

### Declaration of Competing Interest

All authors have completed the ICMJE uniform disclosure form at <http://www.icmje.org/disclosure-of-interest/http://www.icmje.org/disclosure-of-interest/> and declare: All authors had financial support from NordForsk (grant number: 83839) and LG had financial support from Odense University Hospital Research fund (A-number: 2519) and the faculty scholarship from the University of Southern Denmark for the submitted work; Finnish Institute for Health and Welfare (THL) has conducted Public-Private Partnership with vaccine manufacturers and has received research funding from Sanofi Inc., Pfizer Inc., and GlaxoSmithKline Biologicals SA. HN, ML, and AAP have been investigators in these studies, but they have received no personal remuneration; no other relationships or activities that could appear to have influenced the submitted work. Department of Clinical Epidemiology, Aarhus University and Aarhus University Hospital, receives institutional research funding from public and private entities for studies of medicines and vaccines, to and administered by Aarhus University. None of these are relevant to the current study. SS is a salaried employee of Department of Clinical Epidemiology.

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NA.

### Appendix A. Supporting information

Supplementary data associated with this article can be found in the online version at doi:[10.1016/j.jinf.2024.106365](https://doi.org/10.1016/j.jinf.2024.106365).

### References

- Benn CS, Fisker AB, Rieckmann A, Sørup S, Aaby P. Vaccinology: time to change the paradigm? *Lancet Infect Dis* 2020;**20**(10):e274–83.
- Aaby P, Benn CS. Developing the concept of beneficial non-specific effect of live vaccines with epidemiological studies. *Clin Microbiol Infect* 2019;**25**(12):1459–67.
- Higgins JP, Soares-Weiser K, López-López JA, Kakourou A, Chaplin K, Christensen H, et al. Association of BCG, DTP, and measles containing vaccines with childhood mortality: systematic review. *BMJ* 2016;**355**:i5170.
- Sørup S, Benn CS, Stensballe LG, Aaby P, Ravn H. Measles-mumps-rubella vaccination and respiratory syncytial virus-associated hospital contact. *Vaccine* 2015;**33**(1):237–45.
- Sørup S, Benn CS, Poulsen A, Krause TG, Aaby P, Ravn H. Live vaccine against measles, mumps, and rubella and the risk of hospital admissions for nontargeted infections. *JAMA* 2014;**311**(8):826–35.
- Bardenheier BH, McNeil MM, Wodi AP, McNicholl J, DeStefano F. Risk of non-targeted infectious disease hospitalizations among U.S. children following inactivated and live vaccines, 2005–2014. *Clin Infect Dis* 2017;**65**:729–37.
- Jensen A, Andersen PK, Stensballe LG. Early childhood vaccination and subsequent mortality or morbidity: are observational studies hampered by residual confounding? *A Danish register-based cohort study. BMJ Open* 2019;**9**(9):e029794.
- Tielemans SMAJ, de Melker HE, Hahné SJM, Boef AGC, van der Klis FRM, Sanders EAM, et al. Non-specific effects of measles, mumps, and rubella (MMR) vaccination in

- high income setting: population based cohort study in the Netherlands. *BMJ* 2017;**358**:j3862.
9. Sinzinger AX, Von Kries R, Siedler A, Wichmann O, Harder T. Non-specific effects of MMR vaccines on infectious disease related hospitalizations during the second year of life in high-income countries: a systematic review and meta-analysis. *Hum Vaccin Immunother* 2020;**16**(3):490–8.
  10. Sørup S. Careful consideration of hypotheses and model assumptions in study of non-specific effects of vaccines. *Vaccine* 2020;**38**(9):2115.
  11. Gehrt L, Aaby P, Benn CS, Sørup S. Early childhood vaccination and subsequent mortality or morbidity: are observational studies hampered by residual confounding? A Danish register-based cohort study. (<https://bmjopen.bmj.com/content/9/9/e029794.responses#regarding-%E2%80%99Early-childhood-vaccination-and-subsequent-mortality-or-morbidity-are-observational-studies-hampered-by-residual-confounding-a-danish-register-based-cohort-study%E2%80%999D2019>). [cited 2021].
  12. Andrews N, Stowe J, Thomas SL, Walker JL, Miller E. The risk of non-specific hospitalised infections following MMR vaccination given with and without inactivated vaccines in the second year of life. Comparative self-controlled case-series study in England. *Vaccine* 2019;**37**(36):5211–7.
  13. Lawlor DA, Tilling K, Davey Smith G. Triangulation in aetiological epidemiology. *Int J Epidemiol* 2016;**45**(6):1866–86.
  14. Benn CS, Fisker AB, Rieckmann A, Jensen AKG, Aaby P. How to evaluate potential non-specific effects of vaccines: the quest for randomized trials or time for triangulation? *Expert Rev Vaccines* 2018;**17**(5):411–20.
  15. Yung CF. Non-specific effects of childhood vaccines. *BMJ* 2016;**355**:i5434.
  16. Laugesen K, Ludvigsson JF, Schmidt M, Gissler M, Valdimarsdottir UA, Lunde A, et al. Nordic Health Registry-based research: a review of health care systems and key registries. *Clin Epidemiol* 2021;**13**:533–54.
  17. Gehrt L, Laake I, Englund H, Nieminen H, Feiring B, Lahdenkari M, et al. Cohort Profile: Childhood morbidity and potential non-specific effects of the childhood vaccination programmes in the Nordic countries (NONSENSE): register-based cohort of children born 1990–2017/2018. *BMJ Open* 2023;**13**(2):e065984.
  18. Krause TG, Jakobsen S, Haarh M, Mølbak KJE. The Danish vaccination register. 2012;**17**(17):20155.
  19. Riise ØR, Laake I, Bergsaker MA, Nøkleby H, Haugen IL, Storsæter J. Monitoring of timely and delayed vaccinations: a nation-wide registry-based study of Norwegian children aged < 2 years. *BMC Pediatr* 2015;**15**:180.
  20. Chrapkowska C, Galanis I, Kark M, Lepp T, Lindstrand A, Roth A, et al. Validation of the new Swedish vaccination register – accuracy and completeness of register data. *Vaccine* 2020;**38**(25):4104–10.
  21. Schmidt M, Schmidt SAJ, Adelborg K, Sundbøll J, Laugesen K, Ehrenstein V, et al. The Danish health care system and epidemiological research: from health care contacts to database records. *Clin Epidemiol* 2019;**11**:563–91.
  22. Saunes IS, Karanikolos M, Sagan A. Norway: health system review. *Health Syst Transit* 2020;**22**(1):1–163.
  23. KELA. Treatment costs in public health care eu-healthcare.fi2020 [updated 2020. 10.15. Available from: (<https://www.eu-healthcare.fi/what-you-pay/costs-of-treatment-in-finland/treatment-costs-in-public-health-care/>)].
  24. Knutson H. Patientavgifter och högkostnadsskydd 2020 [updated 2020–03–10. Available from: (<https://www.1177.se/sa-fungerar-varden/kostnader-och-ersattningar/patientavgifter/>)].
  25. Bakken IJ, Ariansen AMS, Knudsen GP, Johansen KI, Vollset SE. The Norwegian Patient Registry and the Norwegian Registry for Primary Health Care: research potential of two nationwide health-care registries. *Scand J Public Health* 2020;**48**(1):49–55. 140349481985973.
  26. Ludvigsson JF, Andersson E, Ekblom A, Feychting M, Kim JL, Reuterwall C, et al. External review and validation of the Swedish national inpatient register. *BMC Public Health* 2011;**11**(1):450.
  27. National Institute for Health and Welfare, Care Register for Health Care. Available from: (<https://thl.fi/en/web/thlfi-en/statistics/information-on-statistics/register-descriptions/care-register-for-health-care2016>) [updated 25 Feb 2016].
  28. Schmidt M, Schmidt SA, Sandegaard JL, Ehrenstein V, Pedersen L, Sørensen HT. The Danish National Patient Registry: a review of content, data quality, and research potential. *Clin Epidemiol* 2015;**7**:449–90.
  29. Munk-Jørgensen P, Bertelsen A, Dahl AA, Lehtinen K, Lindström E, Tomasson K. Implementation of ICD-10 in the Nordic countries. *Nordic J Psychiatr* 1999;**53**(1):5–9.
  30. Gehrt L, Laake I, Englund H, Nieminen H, Benn CS, Feiring B, et al. Hospital contacts for infectious diseases among children in Denmark, Finland, Norway, and Sweden, 2008–2017. *Clin Epidemiol* 2022;**14**:609–21.
  31. Buchanan AL, Hudgens MG, Cole SR, Lau B, Adimora AA. Worth the weight: using inverse probability weighted Cox models in AIDS research. *AIDS Res Hum Retrovir* 2014;**30**(12):1170–7.
  32. Hernan MA, Robins JM. Causal inference: What If. 14 October 2019. Forthcoming edn Boca Raton: Chapman & Hall/CRC; 2020.
  33. Austin PC, Stuart EA. Moving towards best practice when using inverse probability of treatment weighting (IPTW) using the propensity score to estimate causal treatment effects in observational studies. *Stat Med* 2015;**34**(28):3661–79.
  34. Andersen PK, Gill RD. Cox regression model for counting-processes – a large sample study. *Ann Stat* 1982;**10**(4):1100–20.
  35. Schoenfeld D. Partial residuals for the proportional hazards regression model. *Biometrika* 1982;**69**(1):239–41.
  36. Kontopantelis E, Reeves D. metaan: Random-effects meta-analysis. *Stata J* 2010;**10**(3):395–407.
  37. Remschmidt C, Wichmann O, Harder T. Frequency and impact of confounding by indication and healthy vaccinee bias in observational studies assessing influenza vaccine effectiveness: a systematic review. *BMC Infect Dis* 2015;**15**:429.
  38. MacLehose RF, Ahern TP, Lash TL, Poole C, Greenland S. The importance of making assumptions in bias analysis. *Epidemiology* 2021;**32**(5):617–24.
  39. Suppli CH, Rasmussen M, Valentiner-Branth P, Mølbak K, Krause TG. Written reminders increase vaccine coverage in Danish children – evaluation of a nationwide intervention using The Danish Vaccination Register, 2014 to 2015. *Eur Surveill* 2017;**22**(17):30522.
  40. Palmu AAM, Nieminen H, Lahdenkari M, Palmu AA. A retrospective nationwide register-based study to evaluate the non-specific effects of first MMR vaccination among children in Finland. *Vaccine* 2023;**41**(3):805–11.
  41. Sørup S, Benn CS, Poulsen A, Krause TG, Aaby P, Ravn H. Simultaneous vaccination with MMR and DTaP-IPV-Hib and rate of hospital admissions with any infections: a nationwide register based cohort study. *Vaccine* 2016;**34**(50):6172–80.
  42. Martins CL, Benn CS, Andersen A, Balé C, Schatzl-Buchholzer F, Do VA, et al. A randomized trial of a standard dose of Edmonston-Zagreb measles vaccine given at 4.5 months of age: effect on total hospital admissions. *J Infect Dis* 2014;**209**(11):1731–8.
  43. Benn CS, Sørup S, Aaby P. Re: Non-specific effects of measles, mumps, and rubella (MMR) vaccination in high income setting: population based cohort study in the Netherlands. *BMJ* 2017;**358**:j3862.
  44. Netea MG, Domínguez-Andrés J, Barreiro LB, Chavakis T, Divangahi M, Fuchs E, et al. Defining trained immunity and its role in health and disease. *Nat Rev Immunol* 2020;**20**(6):375–88.
  45. Kandasamy R, Voysey M, McQuaid F, de Nie K, Ryan R, Orr O, et al. Non-specific immunological effects of selected routine childhood immunisations: systematic review. *BMJ* 2016;**355**:i5225.
  46. Pollard AJ, Finn A, Curtis N. Non-specific effects of vaccines: plausible and potentially important, but implications uncertain. *Arch Dis Child* 2017;**102**:1077–81. archdis-child-2015-310282.
  47. Roring RJ, Debisarun PA, Botey-Bataller J, Suen KT, Bulut O, Kilic G, et al. MMR vaccination induces a trained immunity program characterized by functional and metabolic reprogramming of  $\gamma\delta$  T cells. *J Clin Invest* 2024;**134**(7).